

## OPTIMAL EXPERIMENTAL DESIGN AND QUADRATIC OPTIMIZATION

REBECCA HAYCROFT — LUC PRONZATO — HENRY P. WYNN —  
 — ANATOLY A. ZHIGLJAVSKY

ABSTRACT. A well known gradient-type algorithm for solving quadratic optimization problems is the method of Steepest Descent. Here the Steepest Descent algorithm is generalized to a broader family of gradient algorithms, where the step-length  $\gamma_k$  is chosen in accordance with a particular procedure. The asymptotic rate of convergence of this family is studied. To facilitate the investigation we re-write the algorithms in a normalized form which enables us to exploit a link with the theory of optimum experimental design.

### Introduction

The steepest descent algorithm in  $\mathbb{R}^d$  has been shown to be equivalent to a special algorithm for updating measures on the real line, see, e.g., [4]. The connection is that when the steepest descent algorithm is applied to the minimization of the quadratic function

$$f(x) = \frac{1}{2}(Ax, x) - (x, y), \quad (1)$$

where  $(x, y)$  is the inner product, it can be translated to the updating of measures in  $[m, M]$  where

$$m = \inf_{\|x\|=1} (Ax, x), \quad M = \sup_{\|x\|=1} (Ax, x)$$

with  $0 < m < M < \infty$ ;  $m$  and  $M$  are the smallest and largest eigenvalues of  $A$ , respectively. The research has developed from the well known result, due to Akaike [1], that for standard steepest descent the renormalized iterates  $\frac{x_k}{\|x_k\|}$  converge to the two-dimensional space spanned by the eigenvectors corresponding to the eigenvalues  $m$  and  $M$ .

---

2000 Mathematics Subject Classification: Primary 62K05; Secondary 65K05, 37N40.  
 Keywords: gradient algorithms, steepest descent algorithm, rate of convergence, design of experiments, optimality criteria.

Let  $g(x) = Ax - y$  be the gradient of the objective function (1). The steepest descent algorithm is  $x_{k+1} = x_k - \frac{(g_k, g_k)}{(Ag_k, g_k)}g_k$ . Using the notation  $\gamma_k = \frac{(g_k, g_k)}{(Ag_k, g_k)}$ , we write the algorithm as  $x_{k+1} = x_k - \gamma_k g_k$ . This can be rewritten in terms of the gradients as

$$g_{k+1} = g_k - \gamma_k Ag_k . \tag{2}$$

The main objective of the paper is studying the family of algorithms (2) where the step-length  $\gamma_k$  is chosen in a general way. To facilitate this study we first rewrite the algorithm (2) in a different (normalized) form and then make a connection with the theory of optimum experimental design.

### Renormalized version of gradient algorithms

Let us convert (2) into a “renormalized” version. First note that

$$(g_{k+1}, g_{k+1}) = (g_k, g_k) - 2\gamma_k (Ag_k, g_k) + \gamma_k^2 (A^2 g_k, g_k) . \tag{3}$$

Letting  $r_k = \frac{(g_{k+1}, g_{k+1})}{(g_k, g_k)}$  and dividing (3) through by  $(g_k, g_k)$  gives

$$r_k = 1 - 2\gamma_k \frac{(Ag_k, g_k)}{(g_k, g_k)} + \gamma_k^2 \frac{(A^2 g_k, g_k)}{(g_k, g_k)} . \tag{4}$$

The value of  $r_k$  can be considered as a rate of convergence of algorithm (2) at iteration  $k$ . Other rates which are asymptotically equivalent to  $r_k$  can be considered as well, see [4] for a discussion. The asymptotic rate of convergence of the gradient algorithm (2) can be defined as  $R = \lim_{k \rightarrow \infty} (r_1 \dots r_k)^{1/k}$ . Of course, this rate may depend on the initial point  $x_0$  or, equivalently, on  $g_0$ .

To simplify the notation, we need to convert to moments and measures. Since we assume that  $A$  is a positive definite  $d$ -dimensional square matrix, we can assume, without loss of generality, that  $A$  is a diagonal matrix  $\Lambda = \text{diag}(\lambda_1, \dots, \lambda_d)$ ; the elements  $\lambda_1, \dots, \lambda_d$  are the eigenvalues of the original matrix such that  $0 < m = \lambda_1 \leq \dots \leq \lambda_d = M$ . Then for any vector  $g = (g_{(1)}, \dots, g_{(d)})^T$  we can define

$$\mu_\alpha(g) = \frac{(A^\alpha g, g)}{(g, g)} = \frac{(\Lambda^\alpha g, g)}{(g, g)} = \frac{\sum_i g_{(i)}^2 \lambda_i^\alpha}{\sum_i g_{(i)}^2} .$$

This can be considered as the  $\alpha$ th moment of a distribution with mass  $p_i = \frac{g_{(i)}^2}{\sum_j g_{(j)}^2}$

at  $\lambda_i$ ,  $i = 1, \dots, d$ . Using the notation  $\mu_\alpha^{(k)} = \mu_\alpha(g_k)$ , where  $g_k$  are the iterates in (2), we can rewrite (4) as

$$r_k = 1 - 2\gamma_k \mu_1^{(k)} + \gamma_k^2 \mu_2^{(k)} . \tag{5}$$

For the steepest descent algorithm  $\gamma_k$  minimizes  $f(x_k - \gamma g_k)$  over  $\gamma$  and we have  $\gamma_k = \frac{1}{\mu_1^{(k)}}$  and  $r_k = \frac{\mu_2^{(k)}}{\mu_1^{(k)2}} - 1$ . Write  $z_k = \frac{g_k}{\sqrt{(g_k, g_k)}}$  for the normalized gradient and recall that  $p_i = \frac{g_{(i)}^2}{\sum_j g_{(j)}^2}$  is the  $i$ th probability corresponding to a vector  $g$ . The corresponding probabilities for the vectors  $g_k$  and  $g_{k+1}$  are

$$p_i^{(k)} = \frac{(g_k)_{(i)}^2}{(g_k, g_k)} \quad \text{and} \quad p_i^{(k+1)} = \frac{(g_{k+1})_{(i)}^2}{(g_{k+1}, g_{k+1})} \quad \text{for } i = 1, \dots, d.$$

Now we are able to write down the re-normalized version of (2), which is the updating formula for  $p_i$  ( $i = 1, \dots, d$ ):

$$\begin{aligned} p_i^{(k+1)} &= \frac{(1 - \gamma_k \lambda_i)^2}{(g_k, g_k) - 2\gamma_k (Ag_k, g_k) + \gamma_k^2 (A^2 g_k, g_k)} p_i^{(k)} \\ &= \frac{(1 - \gamma_k \lambda_i)^2}{1 - 2\gamma_k \mu_1^{(k)} + \gamma_k^2 \mu_2^{(k)}} p_i^{(k)}. \end{aligned} \tag{6}$$

When two eigenvalues of  $A$  are equal, say  $\lambda_j = \lambda_{j+1}$ , the updating rules for  $p_j^{(k)}$  and  $p_{j+1}^{(k)}$  are identical so that the analysis of the behaviour of the algorithm remains the same when  $p_j^{(k)}$  and  $p_{j+1}^{(k)}$  are confounded. We may thus assume that all eigenvalues of  $A$  are distinct.

## A multiplicative algorithm for optimal design

Optimization in measure spaces covers a variety of areas and optimal experimental design theory is one of them. These areas often introduce algorithms which typically have two features: the measures are re-weighted in some way and the moments play an important role. Both features arise, as we have seen, in the above algorithms; see also [2], [3] and [5] for examples of other algorithms of this type.

In classical optimal design theory for the linear regression model  $y_j = \alpha + \beta x_j + \varepsilon_j$ ,  $x_j \in [m, M]$ , one is interested in functionals of the moment matrix  $M(\xi)$  of a design measure  $\xi$ :

$$M(\xi) = \begin{pmatrix} \mu_0 & \mu_1 \\ \mu_1 & \mu_2 \end{pmatrix}, \tag{7}$$

where  $\mu_\alpha = \mu_\alpha(\xi) = \int x^\alpha d\xi(x)$  are the  $\alpha$ th moments of the measure  $\xi$  and  $\mu_0 = 1$ .

In the theory of optimum design the directional (Fréchet) derivative “towards” a discrete measure  $\xi_x$  mass 1 at a point  $x$  is of importance. This is

$$\frac{\partial}{\partial \alpha} \Phi \left( M[(1 - \alpha)\xi + \alpha\xi_x] \right) \Big|_{\alpha=0} = \text{tr} \left( \overset{\circ}{\Phi}(\xi)M(\xi_x) \right) - \text{tr} \left( \overset{\circ}{\Phi}(\xi)M(\xi) \right), \quad (8)$$

where

$$\overset{\circ}{\Phi}(\xi) = \frac{\partial \Phi}{\partial M} \Big|_{M=M(\xi)} = \begin{pmatrix} \frac{\partial \Phi}{\partial \mu_0} & \frac{1}{2} \frac{\partial \Phi}{\partial \mu_1} \\ \frac{1}{2} \frac{\partial \Phi}{\partial \mu_1} & \frac{\partial \Phi}{\partial \mu_2} \end{pmatrix}.$$

Here  $\Phi$  is a functional on the space of  $2 \times 2$  matrices usually considered as an optimality criterion to be maximized with respect to  $\xi$ . The first term on the right hand side of (8) is

$$\varphi(x, \xi) = \text{tr} \left( \overset{\circ}{\Phi}(\xi)M(\xi_x) \right) = \begin{pmatrix} 1 & x \end{pmatrix} \overset{\circ}{\Phi}(\xi) \begin{pmatrix} 1 \\ x \end{pmatrix} = \frac{\partial \Phi}{\partial \mu_0} + x \frac{\partial \Phi}{\partial \mu_1} + x^2 \frac{\partial \Phi}{\partial \mu_2}.$$

A class of optimal design algorithms is based on the multiplicative updating of the weights of the current design measure  $\xi^{(k)}$  with some function of  $\varphi(x, \xi)$ . We show below how algorithms in this class are related to the gradient algorithms (2) in their re-normalized form (6).

Assume that our measure is discrete and concentrated on  $[m, M]$ . Assume also that  $\frac{\partial \Phi(M)}{\partial \mu_2} > 0$ ; then  $\varphi(x, \xi)$  has a well-defined minimum

$$c(\xi) = \min_{x \in \mathbb{R}} \varphi(x, \xi) = \frac{\partial \Phi}{\partial \mu_0} - B(\xi), \quad \text{where} \quad B(\xi) = \frac{1}{4} \frac{\left( \frac{\partial \Phi}{\partial \mu_1} \right)^2}{\left( \frac{\partial \Phi}{\partial \mu_2} \right)}.$$

Let  $\xi(x)$  be the mass at a point  $x$  and define the re-weighting at  $x$  by

$$\xi'(x) = \frac{\varphi(x, \xi) - c(\xi)}{b(\xi)} \xi(x), \quad (9)$$

where  $b(\xi)$  is a normalizing constant

$$b(\xi) = \int_m^M (\varphi(x, \xi) - c(\xi)) \xi(dx) = \text{tr} \left[ M(\xi) \overset{\circ}{\Phi}(\xi) \right] - c(\xi).$$

Let us define  $\gamma = \gamma(\xi) = \gamma(\mu_1, \mu_2)$  as

$$\gamma = \gamma(\xi) = \frac{-2 \frac{\partial \Phi}{\partial \mu_2}}{\frac{\partial \Phi}{\partial \mu_1}}. \quad (10)$$

Then

$$\varphi(x, \xi) - c(\xi) = B(\xi) (1 - \gamma(\xi)x)^2.$$

The normalization ensures that the measure  $\xi'$  is a probability distribution. We obtain that the re-weighting formula (9) can be equivalently written as

$$\xi'(x) = \frac{(1 - \gamma x)^2}{1 - 2\gamma\mu_1 + \gamma^2\mu_2} \xi(x). \quad (11)$$

This is exactly the same as the general gradient algorithm in its renormalized form (6). To see that, we simply write the updating formula (11) iteratively

$$\xi^{(k+1)}(x) = \frac{(1 - \gamma_k x)^2}{1 - 2\gamma_k \mu_1^{(k)} + \gamma_k^2 \mu_2^{(k)}} \xi^{(k)}(x).$$

### Optimum design gives the worst rate of convergence

Let  $\Phi = \Phi(M(\xi))$  be an optimality criterion, where  $M(\xi)$  is as in (7). Associate with it a gradient algorithm with step-length  $\gamma(\mu_1, \mu_2)$  as given by (10).

Let  $\xi^*$  be the optimum design for  $\Phi$  on  $[m, M]$ ; that is,

$$\Phi(M(\xi^*)) = \max_{\xi} \Phi(M(\xi))$$

where the maximum is taken over all probability measures supported on  $[m, M]$ . Note that  $\xi^*$  is invariant for one iteration of the algorithm (11); that is, if  $\xi = \xi^*$  in (11) then  $\xi'(x) = \xi(x)$  for all  $x \in \text{supp}(\xi)$ .

In accordance with (5), the rate associated with the design measure  $\xi$  is defined by

$$r(\xi) = 1 - 2\gamma\mu_1 + \gamma^2\mu_2 = \frac{b(\xi)}{B(\xi)}.$$

Assume that the optimality criterion  $\Phi$  is such that the optimum design  $\xi^*$  is non-degenerate (that is,  $\xi^*$  is not just supported at a single point). Note that if  $\Phi(M) = -\infty$  for any singular matrix  $M$ , then this condition is satisfied.

Since the design  $\xi^*$  is optimum, all directional derivatives are non-positive:

$$\frac{\partial}{\partial \alpha} \Phi \left[ M((1 - \alpha)\xi^* + \alpha\xi(x)) \right] \Big|_{\alpha=0^+} \leq 0,$$

for all  $x \in [m, M]$ . Using (8), this implies

$$\max_{x \in [m, M]} \varphi(x, \xi^*) \leq t^* = \text{tr} \left[ M(\xi^*) \overset{\circ}{\Phi}(\xi^*) \right].$$

Since  $\varphi(x, \xi^*)$  is a quadratic convex function of  $x$ , this is equivalent to  $\varphi(m, \xi^*) \leq t^*$  and  $\varphi(M, \xi^*) \leq t^*$ . As

$$\int_m^M \varphi(x, \xi^*) \xi^*(dx) = t^*$$

this implies that  $\xi^*$  is supported at  $m$  and  $M$ . Since  $\xi^*$  is non-degenerate,  $\xi^*$  has positive masses at both points  $m$  and  $M$  and

$$\varphi(m, \xi^*) = \varphi(M, \xi^*) = t^*.$$

As  $\varphi(x, \xi^*)$  is quadratic in  $x$  with its minimum at  $\frac{1}{\gamma}$ , this implies that

$$\gamma^* = \gamma(\mu_1(\xi^*), \mu_2(\xi^*)) = \frac{2}{(m + M)}.$$

The rate  $v(\xi^*)$  is therefore

$$v(\xi^*) = \frac{b(\xi^*)}{B(\xi^*)} = \frac{t^* - c(\xi^*)}{B(\xi^*)} = (1 - m\gamma^*)^2 = (1 - M\gamma^*)^2 = R_{\max},$$

where

$$R_{\max} = \frac{(M - m)^2}{(M + m)^2}. \tag{12}$$

Assume now that the optimum design  $\xi^*$  is degenerate and is supported at a single point  $x^*$ . Note that since  $\varphi(x, \xi^*)$  is both quadratic and convex,  $x^*$  is either  $m$  or  $M$ . Since the optimum design is invariant in one iteration of the algorithm (11),  $\gamma^*$  is constant and

$$\max_{\xi} r(\xi) = \max_{\xi} [(1 - m\gamma^*)^2, (1 - M\gamma^*)^2] \geq R_{\max}$$

with the inequality replaced by an equality if and only if  $\gamma^* = \frac{2}{(M+m)}$ .

### Some special cases

A few examples of gradient algorithms (2) worthy of mention are given below.

The steepest descent algorithm corresponds to the case when  $\Phi(\xi)$  is the  $D$ -optimality criterion  $\Phi(M(\xi)) = \mu_2 - \mu_1^2$  with a step length equal to  $\gamma_k = \frac{1}{\mu_1^{(k)}}$ . It is well-known that the asymptotic rate of the steepest descent algorithm is always close to the value  $R_{\max}$  defined in (12). The asymptotic behaviour of the steepest descent algorithm has already been extensively studied, see, e.g., [4].

The steepest descent algorithm with relaxation is also known in literature on optimization. For this algorithm,  $\gamma_k = \frac{\varepsilon}{\mu_1^{(k)}}$ , where  $\varepsilon$  is some fixed positive number. This algorithm can be associated with the optimality criterion

$$\Phi(M(\xi)) = \varepsilon\mu_2 - \mu_1^2. \tag{13}$$

It is known that for suitable values of the relaxation parameter  $\varepsilon$  this algorithm has a faster convergence rate than the ordinary steepest descent algorithm. However, the reasons why this occurs were not previously known.

It can be shown that if the relaxation coefficient  $\varepsilon$  is either small ( $\varepsilon < \frac{4Mm}{(M+m)^2}$ ) or large ( $\varepsilon > 1$ ), then for almost all starting points the algorithm asymptotically behaves as if it has started at the worst possible initial point. Equivalently, for these values of  $\varepsilon$  the sequence of designs converges to the optimal design for the criterion (13).

Moreover, if the relaxation parameter is either too small ( $\varepsilon < \frac{2m}{m+M}$ ) or too large ( $\varepsilon > \frac{2M}{m+M}$ ), then the rate of the steepest descent algorithm with relaxation becomes worse than  $R_{\max}$ , the worst-case rate of the standard steepest descent algorithm. This is related to the fact that for these values of  $\varepsilon$  the optimal design for the criterion (13) is degenerate (that is, it is concentrated at a single point). As a consequence, we also obtain a well-known result that if the value of the relaxation coefficient is either  $\varepsilon < 0$  or  $\varepsilon > 2$ , then the steepest descent algorithm with relaxation diverges.

When  $\frac{4Mm}{(m+M)^2} < \varepsilon \leq 1$  the relaxed steepest descent algorithm does not converge to the optimum design and its renormalized version (6) asymptotically exhibits either cyclic or chaotic behaviour. It is within this range of  $\varepsilon$  that improved asymptotic rates of convergence are observed. The behaviour of the asymptotic rate  $R$  is shown in Fig. 1, where we display the asymptotic rates in the

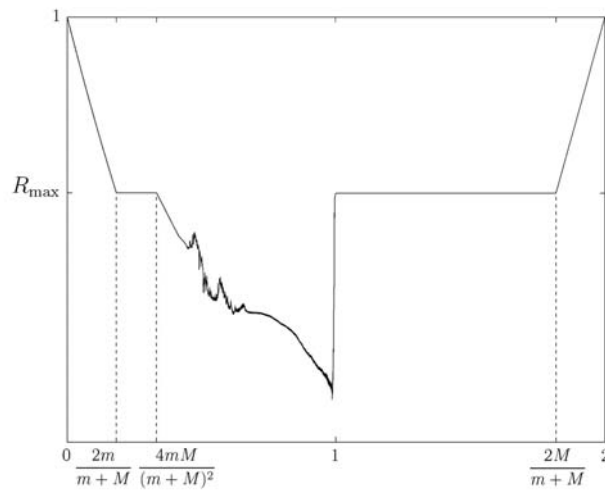


FIGURE 1. Asymptotic rate of convergence as a function of  $\varepsilon$  for the steepest descent algorithm with relaxation  $\varepsilon$ .

case  $\frac{M}{m} = 10$ . In this figure we assume that  $d = 100$  and all the eigenvalues are equally spaced. We have established numerically that the dependence on the dimension  $d$  is insignificant as long as  $d \geq 10$ . In addition, choosing equally spaced

eigenvalues is effectively the same as choosing eigenvalues uniformly distributed on  $[m, M]$  and taking expected values of the asymptotic rates.

The convergence rates of all gradient-type algorithms depend on, amongst other things, the condition number  $\rho = \frac{M}{m}$ . As one would expect, an increase in  $\rho$  gives rise to a worsening rate of convergence. Fig. 2 shows the effect of

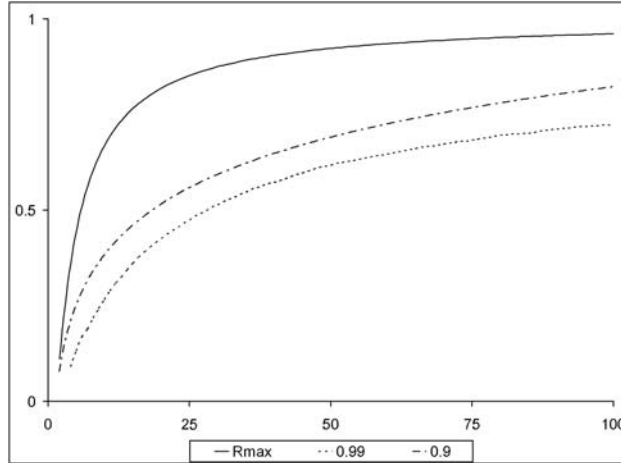


FIGURE 2. Asymptotic rate of convergence as a function of  $\rho$  for steepest descent with relaxation coefficients  $\varepsilon = 0.9$  and  $\varepsilon = 0.99$ .

increasing the value of  $\rho$  on the rates of convergence for the steepest descent algorithm with relaxation coefficients  $\varepsilon = 0.9$  and  $0.99$ .

Any optimization criterion  $\Phi(M(\xi))$  such that  $\frac{\partial \Phi}{\partial \mu_2} > 0$  creates an optimization algorithm of the form (2). Some of these algorithms can be very efficient. For example the family of  $\Phi_p$ -optimality criteria  $\Phi_p(M(\xi)) = (\text{tr} M^{-p}(\xi))^{\frac{1}{p}}$  creates very efficient optimization algorithms. Another useful generalization of the steepest descent algorithm is the family of so-called  $\alpha$ -root algorithms related to the criteria  $\Phi(M(\xi)) = \mu_2^\alpha - \mu_1^{2\alpha}$ . For values of  $\alpha$  slightly larger than 1 the resulting optimization algorithms have been found to be extremely efficient.

#### REFERENCES

- [1] AKAIKE, H.: *On a successive transformation of probability distribution and its application to the analysis of the optimum gradient method*, Ann. Inst. Statist. Math. Tokyo **11** (1959), 1–16.
- [2] MANDAL, S.—TORSNEY, B.: *Construction of optimal designs using a clustering approach*, J. Statist. Plann. Inference **136** (2006), 1120–1134.



- [3] MANDAL, S.—TORSNEY, B.—CARRIERE, K. C.: *Constructing optimal designs with constraints*, J. Statist. Plann. Inference **128** (2005), 609–621.
- [4] PRONZATO, L.—WYNN, H. P.—ZHIGLJAVSKY, A. A.: *Dynamical Search*, Chapman & Hall/CRC, Boca Raton, 2000.
- [5] TORSNEY, B.—MANDAL, S.: *Multiplicative algorithms for constructing optimizing distributions: further developments*. In: Proceedings of the 7th International Workshop on Model-Oriented Design and Analysis—MODA '04 (A. Di Bucchianico et al., eds.), Heeze, The Netherlands, Physica-Verlag, Heidelberg, 2004, pp. 163–171.

Received September 29, 2006

*Rebecca Haycroft*  
*Anatoly A. Zhigljavsky*  
*Cardiff University*  
*School of Mathematics*  
*Senghennydd Road*  
*Cardiff CF24 4AG*  
*UNITED KINGDOM*  
*E-mail: haycroftrj@cf.ac.uk*  
*zhigljavskyaa@cf.ac.uk*

*Luc Pronzato*  
*Laboratoire I3S*  
*2000 route des Lucioles B.P. 121*  
*F-06903 Sophia Antipolis*  
*FRANCE*  
*E-mail: pronzato@i3s.unice.fr*

*Henry P. Wynn*  
*London School of Economics*  
*Dept. of Statistics*  
*London WC2A 2AE*  
*UNITED KINGDOM*  
*E-mail: h.p.wynn@lse.ac.uk*