# BOUNDS FOR DIFFERENTIAL PROBABILITIES IN EVEN ORDER ABELIAN GROUPS 

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#### Abstract

The maximum differential probability for any abelian group of even order is studied. The bounds for these probabilities for groups of order $r$ with $O(\sqrt{r})$ elements of order 2 were given in [T. Tyksiński: Bounds for differential probabilities, Tatra Mt. Math. Publ. 29 (2004), 89-99]. In particular we complete these results by deriving the bounds in the case when the number of elements of order 2 is asymptotically much bigger than the square root of the order of a group.


## 1. Introduction

Differential cryptanalysis is a well known attack on symmetric cryptosystems. It was introduced by Biham and Shamir for DES [2, 3, 4] and it in still serves as a base for many similar attacks, e.g., rectangle attack [1]. Differential cryptanalysis uses pairs of plaintexts with a carefully chosen difference. The primarily used notion of difference was bitwise XOR. A more general definition of difference was introduced later: the difference between two texts from an abelian group $\mathcal{G}=(G, \otimes)$ is defined as $\Delta\left(X, X^{*}\right)=X \otimes\left(X^{*}\right)^{-1}$. Another important structure in differential cryptanalysis is a differential - a pair of differences $(\alpha, \beta)$. These differences as well as texts are elements of $G$; usually $\alpha$ is a difference of plaintexts and $\beta$ is a difference of inputs to the last round of a cipher. Since 1994 Hawkes and O'Connor $[5,6,7,8]$ analysed the behaviour of differentials in commonly used abelian groups and under the assumption of ideal cipher, i.e., a random permutation of texts. We present here new results based on this analysis. First in the next section we introduce the notation and give results obtained by Hawkes, O'Connor [5, 6] and Tyksiński [9]. Then we present the main result of this paper. Finally the last section gives the sketch of proof of the obtained bounds.

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## 2. Bounds in abelian groups

Let $\mathcal{G}=(G, \otimes)$ be an abelian group of order $r$, with a neutral element $e$. Let $\widetilde{\pi}$ be a random permutation selected uniformly from the symmetrical group $S_{r}$. Let us define the random variable $D P_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi})$ describing the number of pairs of difference $\alpha$, that after the permutation $\widetilde{\pi}$ give a difference $\beta$.

$$
D P_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi}):=\left|\left\{\left(X, X \otimes \alpha^{-1}\right) \in G \times G: \Delta\left(\widetilde{\pi}(X), \widetilde{\pi}\left(X \otimes \alpha^{-1}\right)\right)=\beta\right\}\right| .
$$

The probability distribution of $D P_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi})$ is therefore based on the uniform random permutation idealizing the behaviour of a cipher with a randomly chosen key. Moreover we consider the maximum value of $D P_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi})$ defined as

$$
D P_{\mathcal{G}}(\widetilde{\pi}):=\max _{\alpha \neq e, \beta \neq e} D P_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi})
$$

This random variable describes the most probable differential, that can be used in a differential cryptanalysis, based on some definition of difference of texts. Hawkes and $\mathrm{O}^{\prime} \mathrm{Connor}$ wrote a series of papers [5, 6, 7, 8] devoted to the bounds of $D P_{\mathcal{G}}(\widetilde{\pi})$.

The differential probability depends on the notion of difference. The results by O'Connor and Hawkes from [5, 6] apply to XOR and modular addition. Tyksinski in [9] expanded the method and achieved a bound for any abelian group of odd order. Later in [10] Tyksiński tightened a bound in XOR operation abelian groups.
Theorem 1 ([10]). Let $\mathcal{G}=\left(G\right.$, XOR) be an abelian group of order $r=2^{n}$. If $\tilde{\pi}$ is a random permutation (selected uniformly from $S_{r}$ ), then

$$
\operatorname{Pr}\left(\frac{4 \ln r}{\ln \ln r} \leq D P_{\mathcal{G}}(\widetilde{\pi})<\frac{4 \ln r}{\ln \ln r}+\omega(r) \frac{4 \ln \ln \ln r \cdot \ln r}{(\ln \ln r)^{2}}\right) \sim 1
$$

where $\omega(r)$ is any function that goes to infinity arbitrarily slowly as $r \rightarrow \infty$.
An analogous result for abelian groups of odd order was also given in [10]. Proofs were based on the Poisson approximation and tail bounds derived for groups of order $2^{n}$ in $[5,6,7,8]$ and for groups of odd order $r=p_{1}^{k_{1}} p_{2}^{k_{2}} \ldots p_{t}^{k_{t}}$, where $p_{i}$ are odd prime numbers for all $i=1,2, \ldots, t$ in [9].

In this paper we extend the result for abelian groups of odd order to any abelian group of even order $r$ containing $o(\sqrt{r})$ elements of order 2.
Theorem 2. Let $\mathcal{G}=(G, \otimes)$ be any abelian group of even order $r$ containing $o(\sqrt{r})$ elements of order 2. If $\widetilde{\pi}$ is a random permutation (selected uniformly from $S_{r}$ ), then

$$
\operatorname{Pr}\left(\frac{2 \ln r}{\ln \ln r} \leq D P_{\mathcal{G}}(\widetilde{\pi})<\frac{2 \ln r}{\ln \ln r}+\omega(r) \frac{2 \ln \ln \ln r \cdot \ln r}{(\ln \ln r)^{2}}\right) \sim 1
$$

where $\omega(r)$ is any function that goes to infinity arbitrarily slowly as $r \rightarrow \infty$.
Note that, in fact Lemma 5 implies that the lower bound holds for $O(\sqrt{r})$ elements of order 2.

We also extend Theorem 1 to any even order abelian group. The new bounds are additionally expressed in terms of a parameter $d$ - a number of elements of order 2 in the group $\mathcal{G}$.

Theorem 3. Let $\mathcal{G}=(G, \otimes)$ be any abelian group of even order $r$, different from ( $G, \mathrm{XOR}$ ), containing $d$ elements of order 2, where $r=o\left(d^{2}\right)$. If $\widetilde{\pi}$ is a random permutation (selected uniformly from $S_{r}$ ), then

$$
\operatorname{Pr}\left(\frac{4 \ln d}{\ln \ln r}<D P_{\mathcal{G}}(\widetilde{\pi})<\frac{(4+\varepsilon(r)) \ln d}{\ln \ln r}\right) \sim 1
$$

where $\varepsilon(r)$ is any function that goes to 0 as $r \rightarrow \infty$ and such that

$$
\varepsilon(r)>\frac{4 \ln \ln r-4 \ln \ln d+4 \ln \ln \ln r}{\ln \ln d-\ln \ln \ln r} .
$$

Proof of the results is presented in the following section. Note that using the same approach as in the proof of Theorem 2.5 in [10] one can easily get the upper bound in Theorem 2.

## 3. Proof of the main results

### 3.1. Poisson approximation

Graph notation of differences allow us to state the following lemma describing the Poisson approximation for the distribution of random variable $D P_{\mathcal{G}}(\alpha, \beta, \tilde{\pi})$.

Lemma $1([5,6,9])$. Let $\mathcal{G}=(G, \otimes)$ be an abelian group of order $r$, let $t=$ $o(\sqrt[3]{r})$. If ord $\alpha=$ ord $\beta=2$, then

$$
\operatorname{Pr}\left(D P_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi})=2 t\right)=\frac{e^{-1 / 2}}{2^{t} \cdot t!}\left(1+O\left(\frac{t^{3}}{r}\right)\right)
$$

for all other cases (i.e., ord $\alpha \neq 2$ or ord $\beta \neq 2$ ), we have

$$
\operatorname{Pr}\left(D P_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi})=t\right)=\frac{e^{-1}}{t!}\left(1+O\left(\frac{t^{3}}{r}\right)\right)
$$

It follows from Lemma 1 that, in general, the distribution of $D P_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi})$ depends only on the number of elements of order 2 in the group $\mathcal{G}$.

### 3.2. Upper bounds

To show a more precise upper bound in abelian groups of order $r$ we use the Poisson approximation derived in [6] for groups of order $2^{n}$ and in [9] for groups of odd order.

Let $\mathcal{G}=(G, \otimes)$ be an abelian group of order $r$, let $\widetilde{\pi}$ be a random permutation (selected uniformly from $S_{r}$ ). Define an indicator random variable

$$
\Omega_{\mathcal{G}}(\alpha, \beta, \tilde{\pi}, t):= \begin{cases}1 & \text { if } D P_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi})=t \\ 0, & \text { in all other cases }\end{cases}
$$

We are interested in a random variable $\Omega_{\mathcal{G}}(\widetilde{\pi}, t)$, that counts the number of differentials fulfilled by exactly $t$ pairs, i.e.,

$$
\Omega_{\mathcal{G}}(\widetilde{\pi}, t):=\sum_{\alpha, \beta \neq e} \Omega_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi}, t)
$$

Notice that there can be $2^{j}-1$ elements of order 2 in a group for some nonnegative $j$ and therefore the following holds.

Lemma 2. Let $\mathcal{G}$ be an abelian group of even order $r=q \cdot 2^{n}$ ( $q$ is odd) with $2^{j}-1$ elements of order 2 and let $t=o(\sqrt[3]{r})$. Then the expected value of $\Omega_{\mathcal{G}}(\widetilde{\pi}, t)$ is approximated by

$$
\mathbf{E}\left(\Omega_{\mathcal{G}}(\widetilde{\pi}, t)\right) \sim\left(2^{j}-1\right)^{2} \cdot \frac{e^{-1 / 2}}{2^{t / 2}\left(\frac{t}{2}\right)!}+\left(\left(q 2^{n}-1\right)^{2}-\left(2^{j}-1\right)^{2}\right) \cdot \frac{e^{-1}}{t!}
$$

if $t$ is even, and by

$$
\mathbf{E}\left(\Omega_{\mathcal{G}}(\widetilde{\pi}, t)\right) \sim\left(\left(q 2^{n}-1\right)^{2}-\left(2^{j}-1\right)^{2}\right) \frac{e^{-1}}{t!}
$$

if $t$ is odd.
In the case of the group $\mathcal{G}=(G, \mathrm{XOR})$ of order $2^{n}$ the above expectation is equal to

$$
\mathbf{E}\left(\Omega_{\mathcal{G}}(\widetilde{\pi}, 2 t)\right)=\left(2^{n}-1\right)^{2} \cdot \frac{e^{-1 / 2}}{2^{t} \cdot t!}\left(1+O\left(\frac{t^{3}}{2^{n}}\right)\right)
$$

since $t$ can only be even. For odd $t$ the expectation is zero, hence the upper bound of $D P_{\mathcal{G}}(\widetilde{\pi})$ can now be given by the lemma below, which describes the case when a group contains $d$ elements of order 2 and $r=o\left(d^{2}\right)$. As it was already mentioned (see the comment before Theorem 2) an upper bound in the case, where $d=O(\sqrt{r})$, can be shown in the same manner as for groups of odd order (see [10]).

Lemma 3. Let $\mathcal{G}=(G, \otimes)$ be an abelian group of order $r=q \cdot 2^{n}$ ( $q$ is odd), containing d elements of order 2. Moreover let us assume that $r=o\left(d^{2}\right)$. If $\widetilde{\pi}$ is a random permutation (selected uniformly from $S_{r}$ ), then

$$
\operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi})<\frac{(4+\varepsilon(r)) \ln d}{\ln \ln r}\right) \sim 1
$$

where $\varepsilon(r)$ is a function that goes to 0 as $r \rightarrow \infty$, but

$$
\varepsilon(r)>\frac{4 \ln \ln r-4 \ln \ln d+4 \ln \ln \ln r}{\ln \ln d-\ln \ln \ln r} .
$$

Proof. We will show that for

$$
B=\frac{(4+\varepsilon(r)) \ln d}{\ln \ln r}
$$

where $\varepsilon(r)$ is an arbitrarily small positive number, such that

$$
\varepsilon(r)>\frac{4 \ln \ln r-4 \ln \ln d+4 \ln \ln \ln r}{\ln \ln d-\ln \ln \ln r}
$$

we have

$$
\lim _{r \rightarrow \infty} \operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi})>B\right)=0
$$

Let us define a function

$$
k(r):=\left\lfloor\frac{\sqrt[3]{r}}{\bar{\omega}(r)}\right\rfloor
$$

where $\bar{\omega}(r)$ goes to infinity arbitrarily slowly, as $r$ tends to infinity. For the random variable $D P_{\mathcal{G}}(\widetilde{\pi})$ we have

$$
\begin{aligned}
\operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi}) \geq B\right) & =\sum_{t=B}^{k(r)} \operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi})=t\right)+\operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi})>k(r)\right) \\
& \leq \sum_{t=B}^{k(r)} \sum_{\alpha, \beta \neq e} \mathbf{E}\left(\Omega_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi}, t)\right)+\operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi})>k(r)\right) \\
& =\sum_{t=B}^{k(r)} \mathbf{E}\left(\Omega_{\mathcal{G}}(\widetilde{\pi}, t)\right)+\operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi})>k(r)\right)
\end{aligned}
$$

Lemma 2 implies that

$$
\begin{aligned}
& \operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi}) \geq B\right) \leq \sum_{\substack{t=B \\
t \text { is even }}}^{k(r)} \frac{d^{2} \cdot e^{-1 / 2}}{2^{t / 2} \cdot(t / 2)!} \cdot\left(1+O\left(\frac{t^{3}}{r}\right)\right) \\
& \quad+\sum_{t=B}^{k(r)}\left((r-1)^{2}-d^{2}\right) \cdot \frac{e^{-1}}{t!} \cdot\left(1+O\left(\frac{t^{3}}{r}\right)\right)+\operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi})>k(r)\right)
\end{aligned}
$$

From Markov inequality we obtain

$$
\operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi})>k(r)\right) \leq \frac{\mathbf{E}\left(D P_{\mathcal{G}}(\widetilde{\pi})\right)}{k(r)}
$$

Since we estimate probabilities in an abelian group $\mathcal{G}=(G, \otimes)$ containing at most as many elements of order two as in the group $\mathcal{G}^{*}:=(G, \mathrm{XOR})$, therefore by Theorem 3.1 from [8]

$$
\mathbf{E}\left(D P_{\mathcal{G}}(\widetilde{\pi})\right) \leq \mathbf{E}\left(D P_{\mathcal{G}^{*}}(\widetilde{\pi})\right) \leq \frac{2 \ln r}{\ln 2}
$$

Hence

$$
\begin{aligned}
\operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi}) \geq B\right) \leq & d^{2} \sum_{\substack{t=B \\
t \text { is even }}}^{k(r)} \frac{e^{-1 / 2}}{2^{t / 2} \cdot(t / 2)!} \cdot\left(1+O\left(t^{3} / r\right)\right) \\
& +\left((r-1)^{2}-d^{2}\right) \sum_{t=B}^{k(r)} \cdot \frac{e^{-1}}{t!} \cdot\left(1+O\left(t^{3} / r\right)\right)+\frac{2 \ln r}{k(r) \ln 2} .
\end{aligned}
$$

By Stirling's formula we have

$$
\begin{aligned}
\operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi}) \geq B\right) \leq & \frac{d^{2}}{\sqrt{e \pi}} \sum_{\substack{t=B \\
t \text { is even }}}^{k(r)}\left(\frac{e}{t}\right)^{t / 2} \frac{1}{\sqrt{t}} \cdot\left(1+O\left(t^{3} / r\right)\right) \\
& +\frac{(r-1)^{2}-d^{2}}{e \sqrt{2 \pi}} \sum_{t=B}^{k(r)}\left(\frac{e}{t}\right)^{t} \frac{1}{\sqrt{t}} \cdot\left(1+O\left(t^{3} / r\right)\right)+\frac{2 \ln r}{k(r) \ln 2},
\end{aligned}
$$

and therefore we obtain the following bound for the above probability

$$
\left(\frac{d^{2}}{\sqrt{e \pi B}}\left(\frac{e}{B}\right)^{B / 2}+\frac{(r-1)^{2}-d^{2}}{e \sqrt{2 \pi B}}\left(\frac{e}{B}\right)^{B}\right)\left(1+O\left(\frac{k(r)^{3}}{r}\right)\right)+O\left(\frac{\ln r}{k(r)}\right) .
$$

First we will estimate the logarithm of the first summand of this bound
$\ln \frac{(e / B)^{B / 2} \cdot d^{2}}{\sqrt{e \pi B}}=\left(\frac{(4+\varepsilon(r))}{2 \ln \ln r}(1-\ln (4+\varepsilon(r))-\ln \ln d+\ln \ln \ln r)+2\right) \cdot \ln d$

$$
-\frac{1}{2}-\frac{1}{2} \ln \pi-\frac{1}{2} \ln \left(\frac{(4+\varepsilon(r)) \ln d}{\ln \ln r}\right) .
$$

Note that

$$
-\frac{1}{2} \ln \left(\frac{(4+\varepsilon(r)) \ln d}{\ln \ln r}\right) \rightarrow-\infty \quad \text { as } \quad r \rightarrow \infty
$$

Therefore if we show that the coefficient of $\ln d$ is negative, then the whole considered expression would tend to $-\infty$ as $r \rightarrow \infty$. Since $r=o\left(d^{2}\right)$ and $\ln r=$ $o(\sqrt{r})$,

$$
\begin{aligned}
& \frac{(4+\varepsilon(r))}{2 \ln \ln r}-\frac{(4+\varepsilon(r)) \cdot \ln (4+\varepsilon(r))}{2 \ln \ln r}-\frac{(4+\varepsilon(r)) \cdot \ln \ln d}{2 \ln \ln r} \\
& \quad+\frac{(4+\varepsilon(r)) \cdot \ln \ln \ln r}{2 \ln \ln r}+2 \\
& \leq-\frac{(4+\varepsilon(r)) \cdot \ln \ln d}{2 \ln \ln r}+\frac{(4+\varepsilon(r)) \cdot \ln \ln \ln r}{2 \ln \ln r}+2 .
\end{aligned}
$$

Hence if

$$
(4+\varepsilon(r)) \cdot \ln \ln \ln r<(4+\varepsilon(r)) \cdot \ln \ln d-4 \ln \ln r
$$

which is true whenever

$$
\varepsilon(r)>\frac{4 \ln \ln r-4 \ln \ln d+4 \ln \ln \ln r}{\ln \ln d-\ln \ln \ln r}
$$

the coefficient of $\ln d$ is negative and

$$
\lim _{r \rightarrow \infty} \frac{(e / B)^{B / 2} \cdot d^{2}}{\sqrt{e \pi B}}=0
$$

Similarly one can show that

$$
\lim _{r \rightarrow \infty} \frac{(e / B)^{B} \cdot\left((r-1)^{2}-d^{2}\right)}{e \sqrt{2 \pi B}}=0
$$

since $\ln r / k(r)$ tends to 0 as $r \rightarrow \infty$ and the lemma is proven.

### 3.3. Lower bounds

The lower bound for XOR has been calculated by H awkes and O'Connor in [6]. Now we take a closer look at the lower bound for other abelian groups. First we prove the following result.
Lemma 4. Let

$$
B:=\frac{4 \ln d}{\ln \ln r}
$$

If $0 \leq d=d(r)=o(\sqrt{r})$, then

$$
\frac{d^{2} e^{-1 / 2}}{2^{B / 2}(B / 2)!}=o\left(\frac{\left((r-1)^{2}-d^{2}\right) e^{-1}}{B!}\right)
$$

On the other hand if $r=o\left(d^{2}\right)$, then

$$
\frac{\left((r-1)^{2}-d^{2}\right) e^{-1}}{B!}=o\left(\frac{d^{2} e^{-1 / 2}}{2^{B / 2}(B / 2)!}\right)
$$

Proof. Note that by Stirling's formula we have

$$
\begin{equation*}
\frac{d^{2} \cdot e^{-1 / 2}}{2^{B / 2} \cdot(B / 2)!} \cdot \frac{B!}{\left((r-1)^{2}-d^{2}\right) \cdot e^{-1}} \sim \frac{\sqrt{2 e} \cdot d^{2} \cdot(B / e)^{B / 2}}{(r-1)^{2}-d^{2}} . \tag{1}
\end{equation*}
$$

The logarithm of the right side is equal to

$$
\frac{1}{2}+\frac{1}{2} \ln 2+\frac{B}{2} \ln B-\frac{B}{2}-\ln \left(\frac{(r-1)^{2}-d^{2}}{d^{2}}\right) .
$$

Substituting $B:=\frac{4 \ln d}{\ln \ln r}$ we obtain

$$
\begin{aligned}
& \frac{1}{2}+\frac{1}{2} \ln 2+\frac{2 \ln d \cdot \ln 4}{\ln \ln r}+\frac{2 \ln d \cdot \ln \ln d}{\ln \ln r} \\
& \quad-\frac{2 \ln d \cdot \ln \ln \ln r}{\ln \ln r}-\frac{2 \ln d}{\ln \ln r}-\ln \left((r-1)^{2}-d^{2}\right)+2 \ln d .
\end{aligned}
$$

For $d=o(\sqrt{r})$ the last two elements can be estimated by $-\ln r$. The rest can be estimated by $\frac{1}{2}+\frac{1}{2} \ln 2+2 \ln d$. Hence, for such $d$ the logarithm of the right side of (1) tends to $-\infty$. That concludes the first part of our Lemma. Assume that $d=r^{\frac{1+\varepsilon}{2}}$ and note that the logarithm of the right side of (1) is asymptotically equal to

$$
\begin{aligned}
(1+\varepsilon) \ln r\left(\frac{\ln 4}{\ln \ln r}+1-\frac{\ln \ln \ln r}{\ln \ln r}-\frac{1}{\ln \ln r}\right) & -\ln \left(r^{2}\right)+(1+\varepsilon) \ln r \\
& \sim \ln r\left(2 \varepsilon-\frac{(1+\varepsilon) \ln \ln \ln r}{\ln \ln r}\right)
\end{aligned}
$$

and tends to infinity, under the assumption that

$$
\varepsilon>\frac{\ln \ln \ln r}{2 \ln \ln r-\ln \ln \ln r} .
$$

Now we can show two lemmas about lower bound in any abelian group of even order.

Lemma 5. Let $\mathcal{G}=(G, \otimes)$ be an abelian group of order $r=q \cdot 2^{n}$ ( $q$ is odd), containing d elements of order 2, where $0 \leq d=O(\sqrt{r})$. If $\widetilde{\pi}$ is a random permutation (selected uniformly from $S_{r}$ ), then

$$
\operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi})>\frac{2 \ln r}{\ln \ln r}\right) \sim 1
$$

Proof. By Chebychev's inequality for all $B$ we have

$$
\operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi})<B\right) \leq \operatorname{Pr}\left(\Omega_{\mathcal{G}}(\widetilde{\pi}, B)=0\right) \leq \frac{\operatorname{Var}\left(\Omega_{\mathcal{G}}(\widetilde{\pi}, B)\right)}{\mathbf{E}\left(\Omega_{\mathcal{G}}(\widetilde{\pi}, B)\right)^{2}}
$$

Suppose that $B=o(\sqrt[3]{r})$. For such $B$ the square of the expected value of the random variable $\Omega_{\mathcal{G}}(\widetilde{\pi}, B)$ can be approximated in the following way using Lemma 2 . For even $t=o(\sqrt[3]{r})$,

$$
\mathbf{E}\left(\Omega_{\mathcal{G}}(\widetilde{\pi}, t)\right)^{2}=\left(\frac{d^{2} \cdot e^{-1 / 2}}{2^{t / 2}\left(\frac{t}{2}\right)!}+\frac{\left((r-1)^{2}-d^{2}\right) \cdot e^{-1}}{t!}\right)^{2} \cdot\left(1+O\left(\frac{t^{3}}{r}\right)\right)
$$

and for odd $t=o(\sqrt[3]{r})$,

$$
\mathbf{E}\left(\Omega_{\mathcal{G}}(\widetilde{\pi}, t)\right)^{2}=\left(\left((r-1)^{2}-d^{2}\right) \cdot \frac{e^{-1}}{t!}\right)^{2} \cdot\left(1+O\left(\frac{t^{3}}{r}\right)\right)
$$

Since we are interested in the lower bound for the random variable $D P_{\mathcal{G}}(\widetilde{\pi})$ we will use the smaller one, i.e., the case when $t$ is odd. Now for the variance of $\Omega_{\mathcal{G}}(\widetilde{\pi}, B)$ we will need

$$
\begin{aligned}
\mathbf{E}\left(\Omega_{\mathcal{G}}(\widetilde{\pi}, B)^{2}\right)= & \mathbf{E}\left(\left(\sum_{\alpha, \beta \neq e} \Omega_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi}, B)\right)^{2}\right) \\
= & \sum_{\substack{\alpha, \beta \neq e}} \mathbf{E}\left(\Omega_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi}, B)^{2}\right) \\
& +\sum_{\substack{\alpha, \beta, \delta \neq e \\
\delta \neq \beta}} \mathbf{E}\left(\Omega_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi}, B) \Omega_{\mathcal{G}}(\alpha, \delta, \widetilde{\pi}, B)\right) \\
& +\sum_{\substack{\alpha, \beta, \gamma \neq e \\
\gamma \neq \alpha}} \mathbf{E}\left(\Omega_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi}, B) \Omega_{\mathcal{G}}(\gamma, \beta, \widetilde{\pi}, B)\right) \\
& +\sum_{\substack{\alpha, \beta, \gamma, \delta \neq e \\
\gamma \neq \alpha, \delta \neq \beta}} \mathbf{E}\left(\Omega_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi}, B) \Omega_{\mathcal{G}}(\gamma, \delta, \widetilde{\pi}, B)\right) .
\end{aligned}
$$

For the first sum we have

$$
\begin{aligned}
\sum_{\alpha, \beta \neq e} \mathbf{E}\left(\Omega_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi}, B)^{2}\right) & =\sum_{\alpha, \beta \neq e} \mathbf{E}\left(\Omega_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi}, B)\right) \\
& =\sum_{\alpha, \beta \neq e} \operatorname{Pr}\left(D P_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi})=B\right)
\end{aligned}
$$

Let us divide it into two parts

$$
\sum_{\substack{\alpha, \beta \neq e \\ \text { ord } \alpha=\operatorname{ord} \beta=2}} \frac{e^{-1 / 2}}{2^{B / 2} \cdot(B / 2)!}\left(1+O\left(B^{3} / r\right)\right)+\sum_{\substack{\alpha, \beta \neq e \\ \text { ord } \alpha \neq 2 \text { or } \operatorname{ord} \beta \neq 2}} \frac{e^{-1}}{B!}\left(1+O\left(B^{3} / r\right)\right) .
$$

They can be bounded by

$$
(\sqrt{r}-1)^{2} \frac{e^{-1 / 2}}{2^{B / 2} \cdot(B / 2)!}\left(1+O\left(B^{3} / r\right)\right)+\left((r-1)^{2}-(\sqrt{r}-1)^{2}\right) \frac{e^{-1}}{B!}\left(1+O\left(B^{3} / r\right)\right) .
$$

All the other sums we estimate like in $[5,6,9]$ using difference graphs. Let us consider

$$
\mathbf{E}\left(\Omega_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi}, B) \cdot \Omega_{\mathcal{G}}(\alpha, \delta, \widetilde{\pi}, B)\right)
$$

for $B=o(\sqrt[3]{r})$. Now, depending on the orders of each difference, we can have:

- For $\alpha, \beta, \delta$ such that ord $\alpha=\operatorname{ord} \beta=\operatorname{ord} \delta=2$ we have

$$
\mathbf{E}\left(\Omega_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi}, B) \cdot \Omega_{\mathcal{G}}(\alpha, \delta, \widetilde{\pi}, B)\right)=\left(\frac{e^{-1 / 2}}{2^{B / 2} \cdot(B / 2)!}\right)^{2}\left(1+O\left(B^{3} / r\right)\right)
$$

- For $\alpha, \beta, \delta$ such that ord $\alpha=2$ and exactly one of $\beta$ or $\delta$ has order 2 we have

$$
\mathbf{E}\left(\Omega_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi}, B) \cdot \Omega_{\mathcal{G}}(\alpha, \delta, \widetilde{\pi}, B)\right)=\left(\frac{e^{-1 / 2}}{2^{B / 2} \cdot(B / 2)!} \cdot \frac{e^{-1}}{B!}\right)\left(1+O\left(B^{3} / r\right)\right) .
$$

- For $\alpha, \beta, \delta$ such that ord $\alpha=2$ and ord $\beta \neq 2$, ord $\delta \neq 2$ we have

$$
\mathbf{E}\left(\Omega_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi}, B) \cdot \Omega_{\mathcal{G}}(\alpha, \delta, \widetilde{\pi}, B)\right)=\left(\frac{e^{-1}}{B!}\right)^{2}\left(1+O\left(B^{3} / r\right)\right)
$$

- For $\alpha, \beta, \delta$ such that ord $\alpha \neq 2$ we have

$$
\mathbf{E}\left(\Omega_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi}, B) \cdot \Omega_{\mathcal{G}}(\alpha, \delta, \widetilde{\pi}, B)\right)=\left(\frac{e^{-1}}{B!}\right)^{2}\left(1+O\left(B^{3} / r\right)\right)
$$

The same way one can show the approximations for the expectation

$$
\mathbf{E}\left(\Omega_{\mathcal{G}}(\alpha, \beta, \tilde{\pi}, B) \cdot \Omega_{\mathcal{G}}(\gamma, \beta, \tilde{\pi}, B)\right)
$$

Let $\gamma \neq \alpha, \delta \neq \beta$, then

- For $\alpha, \beta, \gamma, \delta$ such that none of them is of order 2 or such that exactly one of them is of order 2 as well as such that exactly two of them are of order 2, either $\alpha$ and $\gamma$ or $\beta$ and $\delta$ or $\alpha$ and $\delta$ or $\beta$ and $\gamma$ we have

$$
\mathbf{E}\left(\Omega_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi}, B) \cdot \Omega_{\mathcal{G}}(\gamma, \delta, \widetilde{\pi}, B)\right)=\left(\frac{e^{-1}}{B!}\right)^{2}\left(1+O\left(B^{3} / r\right)\right)
$$

- For $\alpha, \beta, \gamma, \delta$ such that exactly two of them are of order 2 , either $\alpha$ and $\beta$ or $\gamma$ and $\delta$ or such that exactly three of them are of order 2 we have

$$
\mathbf{E}\left(\Omega_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi}, B) \cdot \Omega_{\mathcal{G}}(\gamma, \delta, \widetilde{\pi}, B)\right)=\left(\frac{e^{-1 / 2}}{2^{B / 2} \cdot(B / 2)!} \cdot \frac{e^{-1}}{B!}\right)\left(1+O\left(B^{3} / r\right)\right) .
$$

- And finally, for $\alpha, \beta, \gamma, \delta$ such that ord $\alpha=$ ord $\beta=$ ord $\gamma=$ ord $\delta=2$ we have

$$
\mathbf{E}\left(\Omega_{\mathcal{G}}(\alpha, \beta, \widetilde{\pi}, B) \cdot \Omega_{\mathcal{G}}(\gamma, \delta, \widetilde{\pi}, B)\right)=\left(\frac{e^{-1 / 2}}{2^{B / 2} \cdot(B / 2)!}\right)^{2}\left(1+O\left(B^{3} / r\right)\right)
$$

All the above calculations hold for $B=o(\sqrt[3]{r})$. Using the following notation

$$
p_{1}:=\frac{e^{-1}}{B!}\left(1+O\left(B^{3} / r\right)\right), \quad p_{2}:=\frac{e^{-1 / 2}}{2^{B / 2}(B / 2)!}\left(1+O\left(B^{3} / r\right)\right),
$$

we get the inequality

$$
\operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi})<B\right) \leq \frac{L}{M}
$$

where $L$-the numerator-is of the form

$$
\begin{align*}
& p_{2}^{2}\left(d^{4}+2 d^{3}-3 d^{2}\right)+p_{1} p_{2}\left(2 d^{2} r^{2}-2 d^{4}-6 d^{2} r+2 d^{3}+4 d^{2}\right) \\
&  \tag{2}\\
& \quad+p_{1}^{2}\left(-r^{2}+d^{2}+2 r-1\right)+p_{2} d^{2}+p_{1}\left(r^{2}-d^{2}-2 r+1\right)
\end{align*}
$$

and the denominator $M$ is equal to

$$
p_{1}^{2}\left(r^{2}-d^{2}\right)^{2}=p_{1}^{2}\left(d^{4}-2 d^{2} r^{2}+r^{4}\right)
$$

Recall that $d^{2}=O(r)$ and let us take

$$
B:=\frac{2 \ln r}{\ln \ln r} .
$$

Let us consider the summands of the sum in the numerator. Notice that

$$
\frac{p_{2}^{2}\left(d^{4}+2 d^{3}-3 d^{2}\right)}{p_{1}^{2}\left(r^{2}-d^{2}\right)^{2}} \sim \frac{p_{2}^{2} \cdot d^{4}}{p_{1}^{2} \cdot r^{4}}
$$

and

$$
\frac{p_{2}^{2} \cdot d^{4}}{p_{1}^{2} \cdot r^{4}} \leq \frac{p_{2}^{2}}{p_{1}^{2} \cdot r^{2}} \sim\left(\frac{e^{-1 / 2}}{2^{B / 2} \cdot(B / 2)!}\right)^{2} \cdot\left(\frac{B!}{e^{-1}}\right)^{2} \cdot \frac{1}{r^{2}} .
$$

By Stirling's formula the right side of the above inequality is asymptotically equal to

$$
\frac{2 e(B / e)^{B}}{r^{2}} .
$$

Therefore, since

$$
\begin{aligned}
\ln \left(\frac{2 e(B / e)^{B}}{r^{2}}\right) & =1+\ln 2+B \ln B-B-2 \ln r \\
& =1+\ln 2+\frac{2 \ln r \cdot \ln 2}{\ln \ln r}-\frac{2 \ln r \cdot \ln \ln \ln r}{\ln \ln r}-\frac{2 \ln r}{\ln \ln r} \rightarrow-\infty
\end{aligned}
$$

as $r \rightarrow \infty$ we obtain that

$$
\frac{p_{2}^{2}\left(d^{4}+2 d^{3}-3 d^{2}\right)}{p_{1}^{2}\left(r^{2}-d^{2}\right)^{2}}=o(1)
$$

Similarly

$$
\begin{aligned}
\frac{p_{1} p_{2}\left(2 d^{2} r^{2}-2 d^{4}-6 d^{2} r+2 d^{3}+4 d^{2}\right)}{p_{1}^{2}\left(r^{2}-d^{2}\right)^{2}} & \sim \frac{p_{2} \cdot 2 d^{2} r^{2}}{p_{1} \cdot r^{4}} \leq \frac{p_{2}}{p_{1} \cdot r}=o(1) \\
\frac{p_{1}^{2}\left(-r^{2}+d^{2}+2 r-1\right)}{p_{1}^{2}\left(r^{2}-d^{2}\right)^{2}} & \sim \frac{-r^{2}+d^{2}}{r^{4}}=o(1)
\end{aligned}
$$

Moreover, we have

$$
\frac{p_{2} d^{2}}{p_{1}^{2}\left(r^{2}-d^{2}\right)^{2}} \leq \frac{p_{2}}{p_{1}^{2} \cdot r^{3}}
$$

and by Stirling's formula

$$
\frac{p_{2}}{p_{1}^{2} \cdot r^{3}} \sim \frac{2 e^{3 / 2}(B / e)^{3 B / 2} \sqrt{\pi B}}{r^{3}}
$$

The logarithm of the right side is equal to

$$
\begin{aligned}
& \ln \left(\frac{2 e^{3 / 2}(B / e)^{3 B / 2} \sqrt{\pi B}}{r^{3}}\right) \\
& =\frac{3}{2}+\ln 2+\frac{3}{2} B \ln B-\frac{3}{2} B+\frac{1}{2} \ln \pi+\frac{1}{2} \ln B-3 \ln r \\
& =\frac{3}{2}+\ln 2+\frac{3 \ln r \cdot \ln 2}{\ln \ln r}-\frac{3 \ln r \cdot \ln \ln \ln r}{\ln \ln r}-\frac{3 \ln r}{\ln \ln r} \\
& \quad+\frac{1}{2} \ln \pi+\frac{1}{2} \ln \left(\frac{2 \ln r}{\ln \ln r}\right)
\end{aligned}
$$

The leading term in the above sum is equal to

$$
-\frac{3 \ln r \cdot \ln \ln \ln r}{\ln \ln r}
$$

which tends to $-\infty$ as $r \rightarrow \infty$. Hence $p_{2} /\left(p_{1}^{2} \cdot r^{3}\right)=o(1)$. Similarly one can show that

$$
\frac{p_{1}\left(r^{2}-d^{2}-2 r+1\right)}{p_{1}^{2}\left(r^{2}-d^{2}\right)^{2}} \sim \frac{1}{p_{1} \cdot r^{2}} \sim \frac{B!}{e^{-1} \cdot r^{2}} \sim \frac{\sqrt{2 \pi B}}{e^{-1} \cdot r^{2}} \cdot\left(\frac{B}{e}\right)^{B}=o(1)
$$

This implies that $\operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi})<B\right)=o(1)$.
The above lemma is used to prove Theorem 2. Our next lemma is needed to prove Theorem 3. In the case of a group that contains $d$ elements of order 2 and $r=o\left(d^{2}\right)$ we prove the following lemma.
Lemma 6. Let $\mathcal{G}=(G, \otimes)$ be an abelian group of order $r=q \cdot 2^{n}$, where $q$ is odd. Furthermore suppose that there are d elements of order 2 in this group, and that $r=o\left(d^{2}\right)$. If $\widetilde{\pi}$ is a random permutation (selected uniformly from $S_{r}$ ), then

$$
\operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi})>\frac{4 \ln d}{\ln \ln r}\right) \sim 1
$$

Proof. We can repeat the reasoning from the previous proof up to the point of defining

$$
p_{1}:=\frac{e^{-1}}{B!}\left(1+O\left(B^{3} / r\right)\right), \quad p_{2}:=\frac{e^{-1 / 2}}{2^{B / 2}(B / 2)!}\left(1+O\left(B^{3} / r\right)\right) .
$$

From Lemma 4 we can see that now $p_{2}$ is of order larger than $p_{1}$. Factoring out $d^{2} p_{2}$ in the variance we get

$$
d^{2} p_{2}+\left((r-1)^{2}-d^{2}\right) p_{1}=d^{2} p_{2}\left(1+\frac{\left((r-1)^{2}-d^{2}\right) p_{1}}{d^{2} p_{2}}\right) \rightarrow d^{2} p_{2}
$$

as $r \rightarrow \infty$. It implies that

$$
\operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi})<B\right) \leq \frac{L}{d^{4} p_{2}^{2}}
$$

The numerator $L$ for $B=\frac{4 \ln d}{\ln \ln r}$ is equal to

$$
\begin{aligned}
L= & p_{2}^{2}\left(2 d^{3}-3 d^{2}\right)+p_{1} p_{2}\left(2 d^{2} r^{2}-6 d^{2} r-2 d^{4}+2 d^{3}+4 d^{2}\right) \\
& +p_{1}^{2}\left(r^{4}-2 d^{2} r^{2}+5 r^{3}+d^{4}+4 d^{2} r-d^{2}-2 r\right) \\
& +p_{1}\left(r^{2}-d^{2}-2 r+1\right)+p_{2} d^{2} .
\end{aligned}
$$

We have

$$
\frac{p_{2}^{2}\left(2 d^{3}-3 d^{2}\right)}{d^{4} p_{2}^{2}} \sim \frac{1}{d}=o(1), \quad \frac{p_{1} p_{2}\left(2 d^{2} r^{2}-6 d^{2} r-2 d^{4}+2 d^{3}+4 d^{2}\right)}{d^{4} p_{2}^{2}} \sim \frac{p_{1} r^{2}}{p_{2} d^{2}}
$$

for $r=o\left(d^{2}\right)$. Moreover,

$$
\ln \left(\frac{p_{1} r^{2}}{p_{2} d^{2}}\right)=2 \ln r-2 \ln d+\frac{B}{2}-\frac{B}{2} \ln B
$$

which for $B=\frac{4 \ln d}{\ln \ln r}$ and $d \geq r^{\frac{1+\varepsilon}{2}}$ is equal to

$$
\left(\frac{1}{\ln \ln r}-\frac{\ln 4}{\ln \ln r}-\frac{\ln \ln d}{\ln \ln r}+\frac{\ln \ln \ln r}{\ln \ln r}-1\right) \cdot 2 \ln d+2 \ln r .
$$

Since $d>\ln r$, we can rewrite the leading terms, for $d=r^{\frac{1+\varepsilon}{2}}$ in the form

$$
2 \ln r-\frac{(1+\varepsilon) \ln r \ln \ln d}{\ln \ln r}-(1+\varepsilon) \ln r .
$$

The limit of this expression is equal to $-\infty$ for any $\varepsilon>0$. Hence we have

$$
\frac{p_{1} r^{2}}{p_{2} d^{2}}=o(1) \quad \text { and, equivalently, } \frac{p_{1}^{2} r^{4}}{p_{2}^{2} d^{4}}=o(1)
$$

Since the other terms are insignificant we obtain

$$
\operatorname{Pr}\left(D P_{\mathcal{G}}(\widetilde{\pi})>B\right)=o(1)
$$

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## REFERENCES

[1] BIHAM, E.-DUNKELMAN, O.-KELLER, N.: The rectangle attack-rectangling the serpent, in: Advances in Cryptology-EUROCRYPT '01 (B. Pfitzmann, ed.), Lecture Notes in Comput. Sci., Vol. 2045, Springer-Verlag, Berlin, 2001, pp. 340-357.
[2] BIHAM, E.-SHAMIR, A.: Differential Cryptanalysis of the Full 16-round DES. Techical Report 708, Technion, Israel Institute of Technology, Haifa, 1991.
[3] BIHAM, E.-SHAMIR, A.: Differential cryptanalysis of the full 16 -round DES, in: Advances in Cryptology-CRYPTO '92 (E. F. Brickell, ed.), Lecture Notes in Comput. Sci., Vol. 740, Springer-Verlag, Berlin, 1993, pp. 487-496.
[4] BIHAM, E.-SHAMIR, A.: Differential Cryptanalysis of the Data Encryption Standard. Springer-Verlag, New York, 1993.
[5] HAWKES, P.-O'CONNOR, L.: XOR and non-XOR differential probabilities, in: Advances in Cryptology-EUROCRYPT '99 (J. Stern, ed.), Lecture Notes in Comput. Sci., Vol. 1592, Springer-Verlag, Berlin, 1999, pp. 272-285.
[6] HAWKES, P.-O'CONNOR, L.: Asymptotic Bounds on Differential Probabilities. Research Report RZ 3018, IBM Research Report, 1998.
[7] O'CONNOR, L.: On the distribution of characteristics in bijective mappings, J. Cryptology 8 (1995), 67-86.
[8] O'CONNOR, L.: On the distribution of characteristics in bijective mappings, in: Advances in Cryptology-EUROCRYPT '93 (T. Helleseth, ed.), Lecture Notes in Comput. Sci., Vol. 765, Springer-Verlag, Berlin, 1994, pp. 360-370.
[9] TYKSIŃSKI, T.: Foundations of differential cryptanalysis in abelian groups, Information Security Proceedings, Lecture Notes in Comput. Sci., Vol. 2851, Springer-Verlag, Berlin, 2003, pp. 280-294.
[10] TYKSIŃSKI, T.: Bounds for differential probabilities, Tatra Mt. Math. Publ. 29 (2004), 89-99.

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